# **HLA and Mate Selection: No Evidence in South Amerindians**

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The factors thought to influence genetic variation at the with respect to HIA, although there are reports comis<br>mans) encompass nearly every facet of population genetic. HIAA in humaning (e.g., see Pollack et al. 1982)<br>man

**Summary** 2<sup>2</sup> types) (for a review, see Brown and Eklund 1994). There have been reports of nonrandom mating (nega-<br>There are studies showing both that males mate prefer-<br>tive-assortative mating) or preference for individuals of entially with females of a different H-2 type (e.g., see<br> served by Potts et al. (1991) could be generated.

**Introduction** There is no published evidence for nonrandom mating

HLA variation has been extensively examined in Received April 2, 1997; accepted for publication June 13, 1997. South Amerindians, both by tissue typing (Bhatia et al. Address for correspondence and reprints: Dr. Philip Hedrick, De-<br>partment of Biology, Arizona State University, Tempe, AZ 85287-<br>1501. E-mail: hedrick@hedricklab.la.asu.edu<br>© 1997 by The American Society of Human Genetics. 0002-9297/97/6103-0008\$02.00 there is evidence supporting substantial selection fa-



voring heterozygotes (or some other type of balancing and Thomson 1983). The observed and expected hetero-<br>selection): deviations from Hardy-Weinberg propor- zygosities are statistically significantly different for selection): deviations from Hardy-Weinberg propor-<br>tions (Black and Salzano 1981), deviations from Mende-<br>lian expectations in families (Black and Salzano 1981). typic frequencies are used in the calculations below (the lian expectations in families (Black and Salzano 1981), and more even distributions of allelic frequencies than complete list of one- and two-locus genotypic frequen-<br>would be expected under neutrality (Hedrick 1983) Be- cies for females and males for the different tribes is av would be expected under neutrality (Hedrick 1983). Be-<br>low we examine the level of HI A sharing in 194 couples able from P.W.H., on request). low we examine the level of HLA sharing in 194 couples, able from P.W.H., on request).<br>all of whom had offspring who were consistent with On the basis of the studies summarized by Parham all of whom had offspring who were consistent with On the basis of the studies summarized by Parham<br>the parental genotypes from 11 remote tribes of South and Ohta (1996), it appears that three of the serotypic the parental genotypes, from 11 remote tribes of South Amerindians, each with  $\geq 10$  couples, to determine alleles at HLA-A—A\*2402, A\*31012, and A\*68012 whether there is a difference from the mating frequencies (here the first two digits indicate the serotype used

The locations of the 11 tribes sampled are given in<br>
figure 1. These tribes are physically quite separated from changementy in South Amerindians (Parham and<br>
figure 1. These tribes are physically quite separated from chro our sample of 15 couples was a majority of the couples Sharing of Alleles in Mating Pairs with children from the sampled village). The estimated The observed sharing of HLA alleles in these pairs, proportion of couples sampled for the other 10 tribes for a given locus, was determined in the following man-

ranged from .30 for the Urubu-Kaabor to .81 for the Arara, with a mean of .52.

These individuals were typed for HLA-A and HLA-B by use of traditional serotyping protocols (e.g., see Black et al. 1980; Black and Salzano 1981). For HLA-A, four different alleles were observed—02, 24, 31, and 68 and, for HLA-B, five different alleles were observed— 05, 15, 35, 39, and 40. Nearly all populations had all alleles—or all alleles except one— at both loci. There is no evidence of serological blank alleles in these samples, for either HLA-A or HLA-B, or in the larger population samples from these populations, suggesting that, if blank alleles were present, they were at quite low frequency. The average observed and expected heterozygosities for the mating pairs over all populations were .764 and .654, respectively, for HLA-A and .732 and .692, respectively, for HLA-B; in other words, there was a substan-**Figure 1** Lower Amazon basin, showing locations of 11 tribal tial degree of polymorphism in the populations, but it groups included in this study. The numbers denote the tribes, given was much lower than the serotypic var was much lower than the serotypic variation observed in the tables. in European or African populations, which may have 15 –18 alleles for HLA-A and 25 –30 alleles for HLA-B and heterozygosities  $> .9$  for both loci (e.g., see Hedrick and Thomson 1983). The observed and expected hetero-

(here the first two digits indicate the serotype used expected under random mating within the tribe. above, and the remaining digits indicate a specific nucle-<br>otide sequence; serotype 68 previously had been desig-**Subjects, Material, and Methods** mated as "28")—are probably homogeneous in these tribes, as they are in other South Amerindian tribes, for Samples and Extent of Genetic Variation<br>The locations of the 11 tribes campled are given in shows heterogeneity in South Amerindians (Parham and

		respect indicat		
No. OF <b>MATINGS</b>	No Alleles	One Allele	Two Alleles	<i>i</i> . The observe males in
19	5/6.63	12/10.47	2/1.89	numbe:
25	3/4.60	17/15.76	5/4.64	an alle
20	0/1.40	15/13.80	5/4.80	case, b
27	6/8.22	14/14.48	7/4.30	
19	6/3.63	9/11.37	4/4.00	within
10	2/1.00	6/6.40	2/2.60	A are
22	0/2.50	19/14.36	3/5.14	HLA-B
11	1/1.27	9/7.27	1/2.45	
15	3/2.40	12/9.47	0/3.13	
13	1/1.00	7/8.31	5/3.69	<b>Results</b>
13 194	2/2.92 29/35.57	8/7.85 128/119.54	3/2.23 37/38.87	Tabl
				<b>OBSERVED/EXPECTED</b> NO. OF MATINGS SHARING

the first individual is the female, the sharing is zero from excess of matings that share two alleles for HLA-B), the mating 02 02  $\times$  24 68, one from 02 24  $\times$  02 68 or 02 opposite of what negative-assortative mating for HLA<br>02  $\times$  02 68, and two from 02 24  $\times$  02 24. The expected would predict. For HLA-A and HLA-B, the observed  $02 \times 02$  68, and two from 02 24  $\times$  02 24. The expected would predict. For HLA-A and HLA-B, the observed sharing of alleles in mating pairs was calculated sepa-<br>numbers of mating that shared no alleles were 29 and sharing of alleles in mating pairs was calculated separately for each population, by use of the observed female 44, respectively, whereas the expected numbers for both and male genotypic frequencies in the mating pairs. This loci were higher—35.57 and 50.55, respectively. There<br>removes any bias that may occur if Hardy-Weinberg was no overall statistical significance between the obremoves any bias that may occur if Hardy-Weinberg genotypic frequencies are used, and it allows for differ- served and expected numbers for either locus, with a ences, in genotypic frequencies, between females and value of  $\chi^2 = 1.90$  (2 df) and  $\chi^2 = 4.14$  (2 df) for HLAmales. Therefore, the expected frequency of a given mating was calculated as  $\Sigma P_{ij}P_{ij,m}$ , where  $P_{ij,f}$  and  $P_{ij,m}$  are the observed frequencies of the genotype with alleles  $i$  **Table 2** and  $j$  in females and males, respectively, in the matings pairs of a given tribe. Each mating pair is then classified Observed and Expected Numbers of Matings That Share or Do Not as sharing no, one, or two alleles, as shown above, and Share Alleles at HLA-B, in 11 Tribes the mating pairs were summed over each of these mating categories. The same approach was extended to two loci where the genotypic frequencies were the observed twolocus female and male genotypic frequencies. The mating pairs were then categorized as sharing zero to four alleles and were summed over the appropriate mating category. We did not examine sharing of two-locus haplotypes, because the two-locus genotypic frequencies give the sum of the two haplotypes in an individual and because we are unable to determine or estimate accurately the haplotypes in families with only one (or a few) offspring and small numbers in the tribal samples.

In addition, the sharing between mates of specific alleles at a locus was also examined; for example, for allele 02 at HLA-A, a female and a male may share allele 02 (e.g., 02 24  $\times$  02 31), the female may have 02 and the male may not (e.g., 02 24  $\times$  24 24), the female may  $\alpha \chi^2 = 4.14$  (df = 2, not significant).

**Table 1** not have 02 and the male may not (e.g.,  $24\ 31 \times 02$ <br> **O2**), or neither may have it. In general, we can indicate Observed and Expected Numbers of Matings That Share or Do Not  $(0, 2)$ , or neither may have it. In general, we can indicate<br>Share Alleles at HLA-A, in 11 Tribes and second letters refer to the female and the male, respectively, and the absence or presence of an overbar indicates that the individual does or does not have allele *i*. The expected values were calculated by use of the observed frequencies of a given allele in females and males in each population, and the expected and observed numbers were summed over tribes. Only tribes in which an allele was present in both sexes were used. In this case, because allelic frequencies must sum to unity within a tribe, three of the allelic comparisons for HLA-A are independent and four allelic comparisons for HLA-B are independent.

Tables 1 and 2 give the observed and expected number of mating pairs that share no, one, or two alleles at  $A^2$  a  $\chi^2$  = 1.90 (df = 2, not significant). HLA-A and HLA-B, for the 11 tribes. Overall, there is no evidence of an excess of observed matings that share no alleles, and, in fact, for both loci there is a slight ner. As examples of mating pairs from HLA-A, where deficiency of matings that share no alleles (and a slight

		<b>OBSERVED/EXPECTED</b> NO. OF MATINGS SHARING			
TRIBE	NO. OF <b>MATINGS</b>	No Alleles	One Allele	Two Alleles	
1. Tiriyo	19	3/5.84	14/11.00	2/2.16	
2. Waiapi	2.5	7/6.72	15/14.80	3/3.48	
3. Apalai	20	8/7.50	11/10.95	1/1.55	
4. Urubu-Kaapor	27	3/10.48	14/13.41	10/3.11	
5. Asurini Trocara	20	7/4.35	9/12.00	4/3.65	
6. Arara	10	1/2.60	8/6.30	1/1.10	
7. Paranana Velho	22	1/1.86	17/14.45	4/5.68	
8. Xikrin	22	3/2.64	5/6.55	3/1.82	
9. Mundurucu	1.5	9/5.07	4/8.07	2/1.87	
10. Karitiana	13	1/0.85	8/9.00	4/3.15	
11. Cinta Larga	13	2/3.15	7/7.38	4/2.46	
Total <sup>a</sup>	194	44/50.55	112/113.54	38/29.91	

## **Table 3**

		<b>OBSERVED/EXPECTED</b> NO. OF MATINGS SHARING					
Tribe	No. OF <b>MATINGS</b>	No Alleles	One Allele	Two Alleles	Three Alleles	Four Alleles	
1. Tiriyo	19	2/2.53	3/6.26	11/7.58	3/2.37	0/0.26	
2. Waiapi	25	1/1.88	5/6.08	16/11.12	1/4.56	2/1.36	
3. Apalai	20	0/0.45	7/5.85	9/9.43	4/4.00	0/0.25	
4. Urubu-Kaapor	27	0/4.00	8/8.52	9/9.96	6/3.85	4/0.67	
5. Asurini Trocara	19	4/1.89	5/3.47	4/7.53	3/4.89	3/1.21	
6. Arara	10	0/0.30	3/2.60	4/4.30	3/2.60	0/0.20	
7. Paranana Velho	22	0/1.27	2/3.32	1.5/7.4.5	3/6.41	2/3.55	
8. Xikrin	11	1/0.36	2/2.55	5/4.91	3/2.73	0/0.45	
9. Mundurucu	15	3/1.19	6/4.31	4/5.81	2/3.06	0/0.62	
10. Karitiana	13	1/.77	0/0.38	7/7.85	1/1.08	4/2.92	
11. Cinta Larga	13	2/1.38	0/3.00	5/4.92	5/2.92	1/.77	
Total <sup>a</sup>	194	14/16.02	41/46.36	89/80.86	34/38.43	16/12.26	

**Observed and Expected Numbers of Matings That Share or Do Not Share Alleles at HLA-A and HLA-B, in 11 Tribes**

<sup>a</sup>  $\chi^2$  = 3.34 (df = 4, not significant).

A and HLA-B, respectively. For HLA-A, observed and they both do not have allele 15 (78 observed vs. 69.95 expected values were very close for all tribes. For HLA- expected). B, the biggest difference (deficiency) was in the Urubu- A large proportion of the excess (4.00 of 8.05) for Kaabor, in which only three matings shared no alleles, both categories is from the Urubu-Kaapor, in which 7 whereas 10.48 were expected, and in which 10 matings matings that shared allele 15 were observed, whereas shared two alleles, whereas only 3.11 were expected. only 3.0 were expected, and in which 16 matings in The Mundurucu showed a somewhat smaller difference which neither partner had allele 15 were observed, (excess) of matings in which no alleles were shared; 9 whereas only 12.0 were expected. There is an excess of were observed, whereas 5.07 were expected. For the homozygotes over heterozygotes in the Urubu-Kaabor other nine tribes, the observed and expected values were (22 homozygotes observed vs. 15.42 expected, the oppo-

To allow determination of whether an effect not seen at a single locus may become apparent when both loci groups that differed in allele frequencies may have been are simultaneously examined, table 3 gives the observed and expected matings for two loci. Again, there is a **Table 4** slight deficiency of matings that share no alleles—14 observed versus 16.02 expected— and a slight excess of **Observed and Expected Numbers of Matings Summed over 11** matings that share all four alleles—16 observed versus **Tribes, Where First** Allele Is That of the Male 12.26 expected. In other words, the trend again is the opposite of that expected for negative-assortative mating. The differences between observed and expected are again not statistically significant ( $\chi^2 = 3.34$ , 4 df).

To allow determination of whether there is nonrandom mating with respect to a given allele, table 4 gives the observed and expected matings for each allele separately. In this case, the observed and expected numbers are again very close to each other, and none of the  $\chi^2$ values are significant (the sample sizes vary by allele because some alleles were missing from some tribes). The largest  $\chi^2$  value is for allele 15 at HLA-B, where there is an excess of observed over expected, both for the category in which the mates share allele 15 (40 ob-<br>served vs. 31.95 expected) and for the category in which

similar for all mating categories.<br>To allow determination of whether an effect not seen unpublished data]), suggesting that two (or more)

	<b>OBSERVED/EXPECTED</b> NO. OF MATINGS				
LOCUS AND <b>ALLELE</b>	i, i	$i,\overline{i}$	$\overline{i}, i$	$\overline{i}, \overline{i}$	$\chi^2$
$HI.A-A:$					
02	72/69.07	42/44.93	42/44.93	38/35.07	.74
24	53/51.43	36/37.57	40/41.57	43/41.43	.24
31	45/47.44	42/38.56	39/37.56	58/60.44	.54
68	14/13.69	17/17.31	23/23.31	86/85.69	.02
$HI.A-B:$					
15	40/31.95	22/30.05	32/40.05	78/69.95	6.74
35	75/73.92	42/43.08	41/42.08	36/34.92	.11
39	22/24.93	37/34.07	26/23.07	72/74.93	1.07
40	26/24.65	33/34.35	34/35.35	78/76.65	.18
51	11/10.54	24/24.46	24/24.46	57/56.45	.04

combined (the Wahlund effect; e.g., see Hedrick 1985). comparison, the estimated proportional differences in If matings also took place separately in two (or more) these four comparisons are .17, .08, .19, and .20, respecseparate groups, then the excess of like  $\times$  like matings tively (in this last comparison, the similar HLA type is could be explained. However, although the Urubu- favored), all of which, except for the second one, Wede-Kaabor are composed of several components, they do kind et al. (1995) found to be statistically significant. not appear to be strongly endogamous, and there is con- Hedrick and Loeschcke (1996), in evaluating the

from 11 different tribes are consistent with random- those used by women to choose a mate for reproduction. mating expectations. In fact, there is an observed defi- A simple estimate of the statistical power to detect

between different H-2 types, found by Yamazaki et al. sample size is larger. (1976), was .57 (a value of 0 indicates no differential We should note that, because of the low polymorselection, and the maximum selection is 1.0; Hedrick phism found in the tribes that we examined, we have a 1994); the level of female-mate selection between differ- similar number of matings that share either no or two ent H-2 types, found by Egid and Brown (1989), was alleles. In more polymorphic populations, such as most .66; and the level of selection necessary to explain the African, Asian, and European populations, a very high deficiency of heterozygotes in the study by Potts et al. proportion of matings would share no alleles, and nearly (1991) was .69 (Hedrick 1992). all the rest would share one allele. Because of the very

ence of female university students for T-shirts worn by there is less statistical power to detect matings related male students who were either dissimilar or similar to to HLA differences in these populations, given the same the females at HLA-A, HLA-B, and HLA-DR (each of sample size, than in our less polymorphic samples. these preference values is the mean of three T-shirt scores of different men). The mean pleasantness scores Marriage Preferences *per female* were 5.61 and 4.66 for dissimilar and similar There is substantial variation, over these tribes, in males, respectively, when the women were not on birth- marriage preferences; for example, there does appear to control pills ( $N = 31$ ) and were 5.03 and 4.60, respec- be a socially dictated preference for marriage between tively, when the women were on birth-control pills ( $N$  cross-cousins—that is, matings with either mother's Å 18). The mean pleasantness scores *per male* were 5.71 brother's child or father's sister's child, which are two when the women were not on birth-control pills ( $N =$  indians. In an ethnographic sample, Murdoch (1957) 38) and were 4.58 and 5.76, respectively, when the lists 77 Central and South American tribes and states women were on birth-control pills ( $N = 23$ ). When these that 12 (16%) preferred cross-cousin marriage. Two of scores are standardized for the highest value in each these tribes, the Apalai and the Mundurucu, are included

favored), all of which, except for the second one, Wede-

siderable exchange among villages; in other words, the study by Wedekind et al. (1995), made several points excess of homozygotes and like  $\times$  like matings is not that are relevant here. Wedekind et al. (1995) had no explainable on the basis of the known anthropological measure of experimental error that could have been obmeasure of experimental error that could have been obinformation and may be just a chance result. tained by presenting the same T-shirt to the same woman more than once. The reversal in preference **Discussion** among women on birth-control pills, depending on whether the per-female or per-male mean is used, is par-Mating frequencies for HLA-A and HLA-B in a sam- ticularly worrisome and was not explained. Finally, the ple that constitutes a large proportion of the total num- preference examined is for a T-shirt odor, a preference ber of matings with progeny from South Amerindians that is probably based on criteria quite different than

ciency of matings that do not share alleles for HLA-A mating based on HLA can be obtained by weighting the and HLA-B, the opposite of what is expected for nega- values expected from random mating (see tables 1 and tive-assortative mating. 2), by 1,  $1 - s/2$ , and  $1 - s$ , for no alleles shared, one allele shared, and two alleles shared, respectively, and Estimate of Selection and Statistical Power then standardizing them to obtain the same total sample Even though we have examined a large proportion of size. For both HLA-A and HLA-B, an *s* value of .45 all the couples with offspring, within all tribes (except gives a  $\chi^2$  value that is significant at the .05 level. In for the Mundurucu), and even though the observed mat- other words, we do not have enough statistical power ings differed from random-mating expectations, in the in these data to detect a selective difference of  $\lt$ .45, direction opposite of negative-assortative mating, we given the negative-assortative mating model above. Thi given the negative-assortative mating model above. This should determine what level of negative-assortative mat- should have allowed us to detect the extent of selection ing we could detect in this sample. As background, let observed in the mouse experiments discussed above but us mention the main cases in which preference has been would not enable us to detect the level of difference shown. First, in mice, the level of male-mate selection observed by Wedekind et al. (1995), even though our

Second, Wedekind et al. (1995) examined the prefer- low numbers of matings that would share two alleles,

cross-cousins— that is, matings with either mother's of the four types of first-cousin matings—in some Amerlists 77 Central and South American tribes and states these tribes, the Apalai and the Mundurucu, are included in our sample. The known genealogies for these tribes variable for different H-2 haplotypes. It is possible that are not deep, but they reveal no evidence of this prefer- the homogeneity of the genetic background in these ence. Either the preference is not rigidly enforced or, in studies may accentuate any differences in the MHC rethese populations, where each member is related to every gion (see below). It is also possible that the MHC region other member through several lines of descent, it can be that differs between haplotypes in these artificially coninterpreted to pose little restriction. This, coupled with structed mouse lines is larger than that in our study of the small proportion of tribes that appear to have such a natural population. In these mouse studies, as well as a preference, makes it unlikely that a substantial effect in the human study by Wedekind et al. (1995), MHC due to personal preference for HLA-dissimilar types differences included both class I and II loci, whereas in would be masked by such cultural preferences, which the present study we have data only on class I loci. It was would increase sharing of HLA in mates. not known or reported in these other studies whether the

over these tribal groups; for example, Kayapo women significant mating effects, but, if the effect is primarily appear to be potent and free to choose any male partner due to class II loci, then we would have have detected and to marry him after they become pregnant, whereas only an effect resulting from linkage disequilibrium bethe Parakana woman appears to be passed in marriage tween class I and II loci. according to rules that favor dominant men. Other tribes In our examination of HLA-A and HLA-B simultaneappear to fall between these extremes. There is no evi- ously, we considered two-locus genotypes, not two-lodence of the arrangement of marriage, as in some other cus haplotypes. As stated above, we did not have either cultures, except for some of the matings between rela- enough progeny from given matings or enough populatives, as discussed above, and these do not appear to be tion data to adequately estimate haplotype frequencies. common in our samples. It is not obvious how mate selection could be based only

tween very close relatives, such as brother-sister, father- genotype (or phenotype) is the only possible unit indaughter, or uncle-niece matings. Matings between close volved in nonrandom mating in a natural population, relatives would presumably result in higher than average and whether it contains particular haplotypes depends sharing of HLA types, and their avoidance would some- on the extent of linkage disequilibrium between loci. In what bias mating frequencies, toward lower than ex- any case, the level of haplotype sharing should be repected HLA sharing, or the appearance of negative-as- flected in the level of sharing at the genotypic level. sortative mating. This effect, in theory, may increase the Furthermore, the examination of sharing of individual frequencies of heterozygotes in very small populations two-locus haplotypes results, even in our rather unpoly- (for discussion, see Markow et al. 1993) and could have morphic sample, in 20 potential two-locus haplotypes, a similar effect on mating frequencies. However, we ac- which, in turn, would make many classes too small for tually observed higher than expected—not lower than satisfactory statistical analysis. expected—HLA sharing, suggesting that, if such an in- It is possible that, in controlled laboratory studies fluence is important in these tribes, other factors must with congenic lines (e.g., see Yamazaki et al. 1976, have masked it. 1988; Egid and Brown 1989), in simplified environ-

some matings thought to share two alleles would share 1995), various genetic, environmental, or cultural faconly one or no alleles, and some matings thought to tors would be eliminated and that MHC would become share one allele would share none. As discussed above, a stronger cue for mate choice. On the other hand, it is within and between South Amerindians tribes, HLA-A often thought that selection in laboratory experiments has very little heterogeneity within serotype, whereas is reduced relative to that present in nature. Although HLA-B appears to have substantially more. Because there may be different and perhaps substantial selective both loci are similarly consistent with random-mating pressures in a laboratory environment, there may in fact proportions in this study, we feel that it is unlikely that be relaxed selection on other factors necessary for surheterogeneity within serotypic class has a substantial vival and reproduction in natural populations. In the impact on mate selection. populations examined here, the individuals chose mates

the genetic background was identical, but mice were in a less natural situation.

In addition, the determination of mates varies greatly two categories of loci contribute differentially to the

In these tribes there is little evidence of matings be- on haplotypes and be independent of genotypes. The

ments such as the seminatural enclosures of Potts et al. Other Considerations (1991), or in human odor-preference trials in which If there is undetected heterogeneity within serotypes, other odors are eliminated or reduced (Wedekind et al. In the laboratory-mouse studies (e.g., see Yamazaki in the way traditional to their tribes, and this presumet al. 1976; Egid and Brown 1989), individuals were ably includes the types of selection that have always homozygous for all loci and were genetically identical existed in these groups. In other words, the mate selecfor all loci—except for the H-2 region, in which the tion that we have observed may be stronger and/or more strains differed. In the study by Potts et al. (1991), again a reflection of natural selective forces than preferences

Hedrick and Black: HLA and Mate Selection 511

controlled situations is overwhelmed by other genetic, Hum Genet  $35:1055-1057$ <br>environmental or cultural factors in a more natural situ-<br>(1985) Genetics of populations. Jones & Bartlett, Bos- $\frac{1}{\text{e}}$  environmental, or cultural factors in a more natural situation, its importance in influencing evolutionary change<br>
may be of minor significance. In light of this, it seems<br>
important to determine the extent of mate choice for<br>
MHC in reproductive pairs in natural populations of<br> man Swiss population examined by Wedekind et al. Hedrick PW, Kim TJ. Genetics of complex polymorphisms: were associated with MHC genotypes; however, they morphology. Cambridge University Press, New York (in  $\mu$ 

presented no information on MHC-genotype similarity<br>between these males and their female mates.<br>There does appear to be strong selection related to<br>HLA—but unrelated to mate selection—acting in these<br>tribes (e.g., see Blac find no evidence of mate choice based on HLA, there Biometrics  $51:1064-1076$ are large selective effects influencing allelic frequencies, Markow T, Hedrick PW, Zuerlein K, Danilovs J, Martin J, genotypic frequencies, and progeny frequencies, from Vyvial T, Armstrong C (1993) HLA polymorphism in the

## **Acknowledgments** pol 59:664–687

Foundation. We appreciate the comments of Jerram Brown, Pollack MS, Wysocki CJ, Beauchamp GK, Braun D Jr, Cal-Amy Eklund, and the reviewers, on an earlier draft of the laway C, Dupont B (1982) Absence of HLA association or

- Alberts SC, Ober C (1993) Genetic variability in the major<br>histocompatibility complex: a review of non-pathogen-me-<br>diated selective mechanisms. Yearbook Phys Anthropol 36:<br>Trends Genet 9:408-412<br>Posenberg LT Cooperman D,
- 7 Rosenberg LT, Cooperman D, Payne R (1983) HLA and mate<br>
Bhatia KK, Black FL, Smith TA, Prasas ML, Koki GN (1995)<br>
Class I HLA antigens in two long-separated populations:<br>
Melanesians and South Amerinds. Am J Phys Anthrop
- 
- Black FL, Salzano FM (1981) Evidence for heterosis in the 263:265–271<br>HLA system. Am J Hum Genet 33:894–899 Wedekind C. Se
- histocompatibility complex: an integrative review. Am Nat [B] 260:245–249<br>143:435–461 Yamazaki K. Beauch
- Egid K, Brown JL (1989) The major histocompatibility com- L, Boyse EA (1988) Familial imprinting determines H-2 seplex and female mating preferences in mice. Anim Behav lective mating preferences. Science 240:1331 –1332 38:548–550 Yamazaki K, Boyse EA, Mike V, Thaler HT, Mathieson BJ,
- Behav 42:693–694 J Exp Med 144:1324–1335
- Obviously, if the odor signal for mating observed in Hedrick PW (1983) Neutrality or selection at HLA? Am J<br>Introlled situations is overwhelmed by other genetic Hum Genet 35:1055–1057
	-
	-
	-
	-
- (1995). In a recent study in pheasants, von Schantz et parasites and maintenance of MHC variation. In: Singh RS, al. (1996) observed that male survival and spur length Krimbas CK (eds) Evolutionary genetics from molecules to<br>were associated with MHC genotypes: however, they morphology. Cambridge University Press, New York (in
	-
	-
	-
- known parents. Havasupai: evidence for balancing selection. Am J Hum Genet 53:943–952
	- Murdock GP (1957) World ethnographic sample. Am Anthro-
	- Parham P, Ohta T (1996) Population biology of antigen pre-P.W.H. appreciates the support of the National Science sentation by MHC class I molecules. Science 272:67-74
- manuscript. **linkage for variations in sensitivity to the odor of androsten**one. Immunogenetics 15:579–589
- Potts WK, Manning CJ, Wakeland EK (1991) Mating patterns **References in** seminatural populations of mice influenced by MHC ge-
	-
	-
	-
- 291–305<br>Black FL, Berman LL, Gabbay Y (1980) HLA antigens in Wittzell H, Goransson G, Grahn M, Persson<br>South American Indians. Tissue Antigens 16:368–376 evidence for the Hamilton-Zuk model. Proc R Soc Lond [B]
- Wedekind C, Seebeck T, Bettens F, Paeke AJ (1995) MHC-Brown JL, Eklund A (1994) Kin recognition and the major dependent mate preferences in humans. Proc R Soc Lond
	- Yamazaki K, Beauchamp GK, Kupniewski D, Bard J, Thomas
- Eklund A, Egid K, Brown JL (1991) The major histocompati- Abbott J, Boyse J, et al (1976) Control of mating preferences bility complex and mating preferences of male mice. Anim in mice by genes in the major histocompatibility complex.